Topics in model fitting statistics: A focus on robust regression and model diagnostic statistics with application to maternal anaemia data in Malawi

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Background

Maternal Anaemia

- Anaemia is a blood disorder characterized by low concentration of haemoglobin (Meena et al., 2019)
- Significant risk factors include education level, body mass index (BMI), wealth index, place of residence, contraception method during pregnancy, water source (Talukder et al., 2022)
- Commonly associated with physiological changes in pregnancy, gravidity, Age, nutritional deficiencies, infection (Malaria, HIV and Hookworm) (Munasinghe & van den Broek, 2006)
- The burden of maternal anaemia is high in sub-Saharan Africa, which derails safe motherhood campaign efforts in the region (Kassebaum et al., 2016)

Goal of the study

Maternal anaemia data modelling gaps and study objective

- Absence/limitations of studies that applied robust regression methods and diagnostic statistics on maternal anaemia data to observe their performance.
- Regression Diagnostic statistics and robust regression methods help to provide better estimates in presence of unusual observations (Ayinde et al., 2015)
- Hence, their application provide comparable quality in the estimates of the risk factors of maternal anaemia
- Thus, there is need to evaluate performance of these methods when applied to the same data

Goal of the study

study objective

- The study compares performance of robust regression methods and diagnostic statistics when applied to both simulation study and maternal anaemia data in Malawi
- 21,935 mothers aged 15-49 years who participated in 2015-16 Malawi Demographic Health Survey (DHS) and had hemoglobin level known were studied

Regression parameter estimation

Multiple linear regression model

• Let Y be the hemoglobin level outcome and X_i, for *i* = 1, 2..., *p* be the explanatory variables, then multiple linear regression model is given by:

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \ldots + \beta_p X_p + \epsilon \tag{1}$$

- where ε is the error term, assumed to be normally distributed, N(0, σ²I_n) and independent
- The ML estimator for β and σ expressed by $\hat{\beta}_N = (X^T X)^{-1} X^T Y$ and $(\hat{\sigma}_N)^2 = \frac{1}{2\sigma^2} \sum (y_i x_i \hat{\beta}_N)^2$ respectively
- Such that the vector of fitted values is represented by:

$$\hat{Y}_N = X(X^T X)^{-1} X^T Y = HY$$
 (2)

Regression parameter estimation

Quantile regression

- Performs better than OLS regression when the data is skewed, it minimizes the median than mean
- For Hemoglobin level (Y) and it's distribution function F(y) = Pr(Y ≤ y), the θ-th, for 0 < θ < 1, quantile is defined as Q(θ) = inf(x : F(Y) ≥ θ)
- Quantile regression model is given by:

$$Q(y_i) = \beta_0(\theta) + \beta_1(\theta)X_{i1} + \beta_2(\theta)X_{i2} + \ldots + \beta_p(\theta)X_{ip}$$
(3)

- Where *i* = 1, ..., *n* and β_j(θ) is estimated by minimizing the problem,helped by R QUANTREG package; ∑ρ_θ(y_i − β₀(θ) − ∑x_{ij}β_jθ)
- Where $\rho_{\theta}(r)$ is the check loss given by $\rho_{\theta}(r) = \theta \max(r, 0) + (1 - r) \max(-1, 0)$

Diagnostic Statistics tools

Outlier and Leverage measures

- At the i-th data point the unstandardized residual is $e_i = y_i \hat{y}$ and $Var(e_i) = \sigma^2(1 h_i)$
- Where $h_i = X_i(X^T X)^{-1}X_i^T$, i-th diagnonal element is interpreted as amount of Leverage or influence exerted by Y_i on \hat{Y}_i , h_i is large if $h_i \ge 2\frac{p}{n}$ where $p = \sum_{i=1}^{n} h_i$
- The standardized (Studentized) residuals, t_i , given by: $t_i = \frac{e_i}{s(1-h_i)^{\frac{1}{2}}}$

• Where
$$s = \left[\frac{\sum_{n=\rho}^{n} e_{i}^{2}}{n-\rho}\right]^{\frac{1}{2}}$$
 is the estimate for σ

Diagnostic Statistics tools

Influence measures

- To examine the effect of the observations on the parameter, Cook's distance, D_i shows the effect of i-th deleted case on all fitted values, $D_i > 1$ is considered influential, $D_i = \frac{(\hat{Y} - \hat{Y}_i)^T (X^T X)(\hat{Y} - \hat{Y}_i)}{ps^2} = \frac{n-p}{p} \frac{h_i}{1-h_i} t_i$
- DFFITS diagnostic combines the information in the leverage h_i , and the Studentized residual e_i , $DFFITS_i$ is considered large if $DFFITS_i \ge 2[\frac{p}{n}]^{\frac{1}{2}}$: $DFFITS_i = \frac{(\hat{Y} - \hat{Y}_i)}{s_i h_i^{\frac{1}{2}}} = [\frac{h_i}{1 - h_i}]^{\frac{1}{2}} t_i$
- DFBETAS measures influence of i-th case on each regression coefficients, b_k , DFBETAS is considered large if is greater than 1 (small data) or $\frac{2}{\sqrt{n}}$ (large data): $DFBETAS_i = \frac{b_k - b_{k(i)}}{\sqrt{MSE_iC_{kk}}}$ where where C_{kk} is the k-th diagonal element of $(X^T X)^{-1}$

Robust Regression statistics

M-Estimators

 The M-estimator's goal is to minimise a function of the errors, ρ rather than the sum of squared errors. The objective function of the M-estimate is:

$$Min\sum_{i=1}^{n}\rho(\frac{e_i}{s}) = Min\sum_{i=1}^{n}\rho(\frac{Y_i - X_i\beta}{s})$$
(4)

- Where s is estimate of scale often formed from linear combination of the residuals
- A reasonable ρ should have the following properties: $\rho(e) \ge 0$, $\rho(0) = 0, \rho(e) = \rho(-e)$, and $\rho(e_i) \ge \rho(e_i^T)$ for $|e_i| = |e_i^T|$
- Minima solution associated with equation (3) is obtained by taking Gauss-Newton iterations, helped by R ROSEPACK package: $\sum_{i=1}^{n} (\phi) (\frac{Y_i - X_i \beta_i}{s}) X_i \text{ where } \phi \text{ is a derivative of } \rho.$

Robust Regression statistics

S-Estimators and MM-Estimators

- S-estimator is defined by minimization of dispersion of residuals: minimize $S(e_1(\theta), ..., e_n(\theta))$, defined as solution of $\frac{1}{n} \sum_{i=1}^{n} \rho(\frac{e_i}{s}) = K$
- Where s(θ) is a type of robust M-Estimate of scale of residuals, K is a constant and ρ(^{e_i}/_s) is the residual function.
- MM-estimators combine the high asymptotic relative efficiency of M-estimators with the high breakdown of class of estimators called S-estimators
- MM-estimator $\hat{\beta}$ defined as a solution to:

$$\sum_{i=1}^{n} x_{ij}(\phi_1) (\frac{y_i - x_i \beta_i}{s_n}) x_i \text{ for } j = 1, 2, ..., p$$
(5)

• Where j=1,2,...,p,
$$\phi_1(\mu) = \frac{\partial \rho_1(\mu)}{\partial \mu}$$

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Robust Regression statistics

LST Estimators

• LTS Estimator minimizes the sum of trimmed squared residuals and is given by:

$$\hat{\beta}_{LTS} = Min \sum_{i=1}^{n} e_i^2 \tag{6}$$

- Such that $e_{(1)}^2 \leq e_{(2)}^2 \dots \leq e_{(n)}^2$ are the ordered squares residuals and h is defined in the range $\frac{n}{2} + 1 \leq h \leq \frac{3n+p+1}{4}$, with n and p being sample size and number of parameters respectively.
- The largest squared residuals are excluded from the summation in this method

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Descriptive statistics of Hemoglobin level and selected covariates

Mean	Median	Std dev	viation	Min	Max	Ske	wness	Kurtosi	s	s	s	s	s	s	s	s	s	s	s	s	s	s	s	s	s	s	s	s	s		
12.53	12.7	1.7	4	2	23	-(0.52	4.85		_	_																_				
	Variable		Categor	у	Frequ	iency	Percenta	age																							
	BMI		Underweight Normal Weight Overweight		1185		5.42																								
					Normal Weight		1095	В	50.09																						
					6733		30.78																								
			Obese		3,001		13.72																								
	Age		15-24		2655		12.10																								
	Age		25-49		1928	D	87.90																								
			No educ	ation	4451		20.29																								
	Highest education level		Primary Secondary		est education level		1395	В	63.63																						
					3269		14.90																								
			Tertiary		257		1.17																								
			Poor		8699		39.66																								
	Wealth Index		Middle		4403		20.07																								
			Rich		8833		40.27																								
	Water source		Unsafe V		3010		13.79																								
	Water Source		Safe Wa	iter	1882	1	86.21																								
	Distance to he	alth centre	Big prob	olem	1213	3	55.31																								
	Distance to in	carer centre	No prob	lem	9802		44.69																								
	Residence		Urban		3440		15.64																								
	Residence		Rural		1849	5	84.32																								
			Modern	method	1256	7	57.29																								
	Contraceptive	use	Traditio	nal method	272		1.24																								
			Non-use	rs	9096		41.47																								
									•	▶ ∢ 🗗	▶ ∢ 🗗 ▶	→ ◆ 🗇 ▶ ◆ 🖹	▶ ▲圖▶ ▲圖▶	▶ ▲圖▶ ▲ 圖▶ ▲	▶ < □ > < ⊇ > < 3	▶ ▲圖▶ ▲圖▶ ▲圖	▶ ▲圖▶ ▲圖▶ ▲圖	▶ ▲圖▶ ▲圖▶ ▲圖	▶ ▲圖▶ ▲臣▶ ▲臣	▶ ▲圖▶ ▲圖▶ ▲團	→ ★圖 ▶ ★ 臣 ▶ ★ 臣	▶ < @ ▶ < E ▶ < E	▶ < @ ▶ < E ▶ < E	▶ ▲圖▶ ▲ 圖▶ ▲ 圖	▶ ▲圖▶ ▲圖▶ ▲圖	▶ ▲圖▶ ▲ 圖▶ ▲ 圖	▶ ▲圖▶ ▲圖▶ ▲圖!	▶ ▲圖▶ ▲圖▶ ▲圖▶	→ ▲圖 ▶ ▲ 臣 ▶ ▲ 臣 ▶		

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Model Estimates

Multiple Linear and Quantile Regression ML Estimates

Variable	Category	OLS p-value	Q ₂₅ p-value	Q ₅₀ p-value	Q ₇₅ p-value	Q ₉₀ p-value
Intercept		11.86	10.97	12.59	13.46	14.1
		(< 0.0001)	(< 0.0001)	(< 0.0001)	(< 0.0001)	(< 0.0001)
Residence	Urban*					
Residence	Rural	-0.54 (0.001)	-0.32 (0.119)	-0.53 (0.004)	-0.92	-1 (0.001)
					(< 0.0001)	
	No education*					
Education	Primary	0.84	0.59	0.82	1.1	1.40
Education		(< 0.0001)	(< 0.0001)	(< 0.0001)	(< 0.0001)	(< 0.0001)
	Secondary	0.65	0.32 (0.160)	1.01	0.64 (0.015)	0.80 (0.013)
		(< 0.0001)		(< 0.0001)		
	Tertiary	0.33 (0.546)	-0.36 (0.599)	-0.16 (0.799)	1.01 (0.205)	1.10 (0.256)
Gravidity		0.08 (0.003)	0.09 (0.005)	0.08 (0.005)	-0.002	-2.5e ⁷ (1.00)
					(0.950)	
Current pregnant dura-		-0.02	-0.24	-0.02	-0.024	-0.20
tion		(< 0.0001)	(< 0.0001)	(< 0.0001)	(< 0.0001)	(< 0.0001)
Distance to Health Centre	Big problem*					
Distance to Health Centre	No problem	0.2 (0.043)	0.11 (0.367)	0.06 (0.560)	0.10 (0.492)	8.86e ⁷ (1.00)
	Underweight*					
BMI	Normal	0.15 (0.693)	0.15 (0.752)	-0.27 (0.519)	0.15 (0.785)	0.60 (0.366)
Divit	Overweight	0.19 (0.616)	0.28 (0.542)	-0.11 (0.790)	0.31 (0.565)	1.00 (0.134)
	Obese	0.52 (0.199)	0.71 (0.153)	0.03 (0.955)	0.69 (0.235)	1.00 (0.158)
	Poor*					
Wealth Index	Middle	0.20 (0.128)	0.32 (0.048)	0.17 (0.228)	0.43 (0.022)	-0.30 (0.188)
	Rich	-0.20 (0.100)	-0.25 (0.098)	-0.48	-0.15 (0.387)	-0.7 (0.001)
				(< 0.0001)		
Age	15-24*					
	25-49	-0.10 (0.495)	-0.09 (0.615)	-0.45 (0.004)	-0.11 (0.585)	-0.10 (0.683)
Water source	Unsafe*					
	Safe	-0.24 (0.064)	-0.13 (0.393)	-0.26 (0.069)	-0.25 (0.164)	-0.80
						(< 0.0001)
AIC		4352.5	4459.0	4297.5	4487.4	4897.6

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Outliers in the models

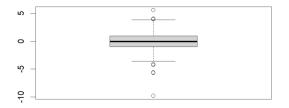


Figure: Box plot for OLS regression model

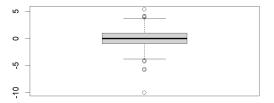


Figure: Box plot for Quantile regression model

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Work in progress

- Apply simulation techniques to compare efficiency of estimates from maximum likelihood estimation and robust regression for a linear model
- To apply simulation methods to compare effectiveness of robust regression and diagnostic statistics in detecting outliers and influential data points to a linear model
- To apply diagnostic statistics and robust regression on maternal anaemia data in Malawi to compare the extent of detection of outliers and influential observations by each method

Conclusion

- The study observes performance of robust regression methods and diagnostic statistics on quality of estimates and detection of unusual observations when applied to both simulated data and maternal anaemia data in Malawi
- Akaike Information Criterion (*AIC*) was used to determine the best model fit for the dataset and p-values were used to determine statistical significance
- 50th Quantile regression model was the best fitted model to the dataset as it has small AIC compared to 25th, 75th, 90th and OLS
- Quantile regression model identified 8 outliers while OLS regression model identified 6 outliers in the dataset

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